

Package ‘nsROC’

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Description Tools for estimating Receiver Operating Characteristic (ROC) curves, building confidence bands, comparing several curves both for dependent and independent data, estimating the cumulative-dynamic ROC curve in presence of censored data, and performing meta-analysis studies, among others.

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nsROC-package

nsROC

Description

Tools for estimating Receiver Operating Characteristic (ROC) curves, building confidence bands, comparing several curves both for dependent and independent data, estimating the cumulative-dynamic ROC curve in presence of censored data, and performing meta-analysis studies, among others.

Details

The basic function of the nsROC package is the [gROC](#) function. It will estimate an ROC curve under one of these considerations: larger values of the marker are associated with a higher probability of being positive (right-sided), the opposite (left-sided) or when both smaller and larger values of the marker are associated with having more probability of being positive (both).

Confidence bands for an ROC curve estimate resulting of the previous function can be computed and displayed by the [ROCbands](#) function. Three different methods are provided to compute them.

Several paired or unpaired ROC curves can be compared with the [compareROCdep](#) or [compareROCindep](#) function, respectively. In order to compare ROC curves different statistics can be used, and to approximate the distribution of the statistic in the paired case both permutation and bootstrap procedures are computed.

Time-dependent ROC curves can be estimated by the cumulative/dynamic approach using the [cdROC](#) function. In order to deal with the right censored problem three different statistics can be considered.

Meta-analysis of ROC curves following a non-parametric approach can be performed with the [metaROC](#) function. Both the fixed-effects and random-effects model can be considered.

Abbreviations

The following abbreviations are frequently used in this package:

- ROC: Receiver Operating Characteristic
- AUC: Area Under the (ROC) Curve
- Sp: Specificity
- Se: Sensitivity
- TPR: True-Positive Rate
- FPR: False-Positive Rate

Functions

gROC	ROC curve estimate (generalization included)
ROCbands	Confidence bands for ROC curves
compareROCdep	Comparison of k paired ROC curves
compareROCindep	Comparison of k independent ROC curves
cdROC	Cumulative/dynamic ROC curve estimate
metaROC	Non-parametric ROC curve estimate for meta-analysis
plot	Plot an ROC curve
plot	Plot confidence bands for an ROC curve
plot	Plot a time-dependent ROC curve
print	Print a groc object
print	Print a rocbands object
print	Print a cdroc object
checkROC	Check the data to compute an ROC curve (internal function)

Dataset

This package comes with a dataset of 9 papers (meta-analysis) with the number of TP (true positive), FP (false positive), TN (true negative) and FN (false negative) about the use of the Interleukin6 (IL6) as a marker for the early detection of neonatal sepsis: [interleukin6](#).

Installing and using

To install this package:

```
install.packages("nsROC")
```

To load the package:

```
library(nsROC)
```

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See Also

CRAN packages **sde** and **survival** employed in this package.

Description

This function estimates a time-dependent ROC curve following the cumulative/dynamic approach and returns a 'cdroc' object. This object can be printed or plotted. To deal with the right censored problem different statistics can be considered: those ones proposed by *Martinez-Cambor et al. (2016)* based on the hazard Cox regression model (semiparametric) or the Kaplan-Meier estimator (non-parametric); and the one included in *Li et al. (2016)* based on the kernel-weighted Kaplan-Meier method. See *References* below.

Usage

```
cdROC(stime, status, marker, predict.time, ...)
## Default S3 method:
cdROC(stime, status, marker, predict.time, method=c('Cox', 'KM', 'wKM'),
       kernel=c('normal', 'Epanechnikov', 'other'), h=1,
       kernel.fun = function(x,xi,h){u <- (x-xi)/h; 1/(2*h)*(abs(u) <= 1)},
       ci=FALSE, boot.n=100, conf.level=0.95, seed=2032, ...)
```

Arguments

stime	vector of observed times.
status	vector of status (takes the value 0 if the subject is censored and 1 otherwise).
marker	vector of (bio)marker values.
predict.time	considered time point (scalar).
method	procedure used to estimate the probability. One of "Cox" (method based on Cox regression), "KM" (method based on Kaplan-Meier estimator) or "wKM" (method based on kernel-weighted Kaplan-Meier estimator).
kernel	procedure used to calculate the kernel function. One of "normal", "Epanechnikov" or "other". Only considered if method='wKM'.
h	bandwidth used to calculate the kernel function. Only considered if method='wKM'.
kernel.fun	if method='wKM' and kernel='other', function used to calculate the kernel function. It has three input parameters: x=vector, xi=value around which the kernel weight should be computed, h=bandwidth. Default: Uniform kernel.
ci	if TRUE, a confidence interval for the area under the curve is computed.
boot.n	number of bootstrap replicates considered to build the confidence interval. Default: 100.
conf.level	the width of the confidence band as a number in (0,1). Default: 0.95, resulting in a 95% confidence band.
seed	seed considered to generate bootstrap replicates (for reproducibility).
...	additional arguments for cdROC. Ignored.

Details

Assuming that larger values of the marker are associated with higher probabilities of occurrence of the event, the cumulative sensitivity and the dynamic specificity are defined by:

$$Se^C(x, t) = P(\text{marker} > x | \text{stime} \leq t) \text{ and } Sp^D(x, t) = P(\text{marker} \leq x | \text{stime} > t).$$

The resulting ROC curve is known as the cumulative/dynamic ROC curve, $R_t^{C/D}$, where $t = \text{predict.time}$.

Data censored before t is the major handicap with regard to the estimation of the time-dependent ROC curve. In order to estimate the probability of surviving beyond t for the i -th subject, \hat{P}_i , three different methods are considered:

- A semiparametric one, using a proportional hazard Cox regression model:

The hazard function is estimated by $\lambda(t) = \lambda_0(t) \cdot \exp(\beta \cdot X)$ where X denotes the marker.

The probability is estimated by $\hat{P}_i = \frac{\hat{S}(t|X=x_i)}{\hat{S}(z_i|X=x_i)}$ where z_i stands for the observed time of the i -th subject and \hat{S} is the survival function estimated from the Cox regression model.

- A non-parametric one, using the Kaplan-Meier estimator directly:

The probability is estimated by $\hat{P}_i = \frac{\hat{S}(t)}{\hat{S}(z_i)}$ where z_i stands for the observed time of the i -th subject and \hat{S} is the survival function estimated by the Kaplan-Meier method referred to those subjects satisfying $X \leq x_i$.

- A non-parametric one, using the kernel-weighted Kaplan-Meier estimator:

The survival function is estimated by $\hat{S}(t|X = x_i) = \prod_{s \leq t} \left[1 - \frac{\sum_{j=1}^n K_h(x_j, x_i) I(z_j = s) \text{status}_j}{\sum_{j=1}^n K_h(x_j, x_i) I(z_j = s)} \right]$ where z_j stands for the observed time of the j -th subject, I is the indicator function and status_j takes the value 0 if the j -th subject is censored and 1 otherwise.

Two different methods can be considered in order to define the kernel function, $K_h(x_j, x_i)$:

– kernel='normal':

$$K_h(x_j, x_i) = \frac{1}{h\sqrt{2\pi}} \exp\left\{-\frac{(x_j - x_i)^2}{2h^2}\right\}$$

– kernel='Epanechnikov':

$$K_h(x_j, x_i) = \frac{3}{4h} \left(1 - \frac{x_j - x_i}{h}\right) I(|x_j - x_i| \leq h)$$

where h is the bandwidth considered for kernel weights.

If the user decide to use another kernel function, kernel='other', it should be defined by the kernel.fun input parameter, which has three parameters following this order: x is a vector, x_i is the value around which the kernel weight should be computed and h is the bandwidth.

The probability is estimated by $\hat{P}_i = \frac{\hat{S}(t|X=x_i)}{\hat{S}(z_i|X=x_i)}$ where z_i stands for the observed time of the i -th subject and \hat{S} is the survival function estimated by the kernel-weighted Kaplan-Meier method considered above.

Value

A list of class 'cdroc' with the following content:

TP	vector of sensitivities (true positive rates).
TN	vector of specificities (true negative rates).

cutPoints	vector of thresholds considered for the (bio)marker. It coincides with the marker vector adding $\min(\text{marker}) - 1$ and $\max(\text{marker}) + 1$.
auc	area under the curve estimate by trapezoidal rule.
ci	if TRUE, a confidence interval for the area under the curve has been computed.
boot.n	number of bootstrap replicates considered to build the confidence interval. Default: 100.
conf.level	the width of the confidence band as a number in (0,1). Default: 0.95, resulting in a 95% confidence band.
seed	seed considered to generate bootstrap replicates (for reproducibility).
meanAuc	bootstrap area under the curve estimate (mean along bootstrap replicates).
ciAuc	bootstrap confidence interval for the area under the curve.
aucs	vector of bootstrap area under the curve estimates.
stime	vector of observed times.
status	vector of status (takes the value 0 if the subject is censored and 1 otherwise).
marker	vector of (bio)marker values.
predict.time	considered time point (scalar).
method	procedure used in order to estimate the probability.
kernel	procedure used to calculate the kernel function. Only considered if method='wKM'.
h	bandwidth used to calculate the kernel function. Only considered if method='wKM'.

Note

survfit and Surv functions in survival package are used in order to estimate the survival functions in both methodologies. Additionally, coxph from the same package is used to fit the Cox proportional hazard regression model in the semiparametric approach.

References

- Martinez-Camblor P., F-Bayon G., Perez-Fernandez S., 2016, Cumulative/dynamic ROC curve estimation, *Journal of Statistical Computation and Simulation*, **86**(17), 3582-3594.
- Li L., Greene T., Hu B., 2016, A simple method to estimate the time-dependent receiver operating characteristic curve and the area under the curve with right censored data, *Statistical Methods in Medical Research*, DOI: 10.1177/0962280216680239.

Examples

```
# Basic example. Data
set.seed(123)
stime <- rchisq(50,3)
status <- sample(c(rep(1,40), rep(0,10)))
marker <- max(stime) - stime + rnorm(50,0,2)

# Cumulative/dynamic ROC curve estimate at time 2.8 (Cox method is used) with 0.95 confidence
# interval for the area under the curve
cdROC(stime, status, marker, 2.8, ci=TRUE)
```

```
# Cumulative/dynamic ROC curve estimate at time 3.1 (Kaplan-Meier method is used)
cdROC(stime, status, marker, 3.1, method="KM")

# Cumulative/dynamic ROC curve estimate at time 3 (kernel-weighted Kaplan-Meier method with
# gaussian kernel and bandwidth 1 is used)
cdROC(stime, status, marker, 3, method="wKM")

# Cumulative/dynamic ROC curve estimate at time 3 (kernel-weighted Kaplan-Meier method with
# biweight kernel and bandwidth equals to 2 is used)
cdROC(stime, status, marker, 3, method="wKM", kernel="other", h=2,
      kernel.fun = function(x,xi,h){u <- (x-xi)/h; 15/(16*h)*(1-u^2)^2*(abs(u)<=1)})
```

checkROC

Check data to compute an ROC curve

Description

This internal function checks if the data introduced for building the curve is correct or not. It shows if there are some missing marker or response values and whether there are less or more than two levels at the response vector, D. It also splits the data into two groups: controls and cases, depending on the corresponding value in the response vector.

Usage

```
checkROC(X,D)
```

Arguments

X vector of (bio)marker values. It should be numeric.
D vector of response values. It should contain at least two different levels.

Details

The code will not run and an error will be showed in these cases:

- X or D is missing,
- X or D is NULL or full of NA's,
- X is not a numeric vector,
- D has less than two different values, and/or
- X and D have different lengths.

If the response vector has more than two different levels, only the two first ones are considered as controls and cases, respectively.

If the user does not agree with the codification, it can be changed modifying the order of the levels using the factor function; for instance, `factor(D, levels=c("1", "0"))`.

Value

If the marker and response vectors are correct a list with the following fields is returned:

levels	levels in D. The two first ones are the labels of D considered as controls and cases, respectively.
controls	marker values for controls.
cases	marker values for cases.
n0	number of controls.
n1	number of cases.
X	marker values corresponding to controls and cases (in this order).
D	response vector consisting of n0 repetitions of levels[1] and n1 repetitions of levels[2] (in this order).

Examples

```
# Basic example with full information
set.seed(123)
X <- c(rnorm(45), rnorm(30,2,1.5))
D <- c(rep(0,45), rep(1,30))
checkROC(X,D)

# Example with some missing values and more than two levels
X <- replace(c(rnorm(25), rnorm(30,2,1.5), rnorm(20,-3,1)), seq(1,75,5), NA)
D <- replace(c(rep(0,25), rep(1,30), rep(2,20)), seq(1,75,11), NA)
checkROC(X,D)
```

compareROCdep

Comparison of k paired ROC curves

Description

This function compares k ROC curves from dependent data. Different statistics can be considered in order to perform the comparison: those ones included in *Martinez-Cambor et al. (2013)* based on general distances between functions, the *Venkatraman et al. (1996)* methodology for comparing diagnostic the accuracy of the k markers based on data from a paired design and the *DeLong et al. (1988)* one based on the AUC (area under the curve) comparison. Two different methods could be considered to approximate the distribution function of the statistic: the procedure proposed by *Venkatraman et al. (1996)* (based on permutated samples) or the one introduced by *Martinez-Cambor et al. (2012)* (based on bootstrap samples). See *References* below.

Usage

```
compareROCdep(X, D, ...)
## Default S3 method:
compareROCdep(X, D, method=c("general.bootstrap", "permutation", "auc"),
              statistic=c("KS", "L1", "L2", "CR", "VK", "other"),
```



```

FUN.dist=function(g){max(abs(g))}, side=c("right","left"),
Ni=1000, B=500, perm=500, seed=123, h.fun=function(H,x){
H*sd(x)*length(x)^{-1/3}}, H=1, plot.roc=TRUE, type='s', lwd=3,
lwd.curves=rep(2,ncol(X)), lty=1, lty.curves=rep(1,ncol(X)),
col='black',col.curves=rainbow(ncol(X)), cex.lab=1.2,
legend=c(sapply(1:ncol(X), function(i){eval(bquote(expression(
hat(R)[.(i)](t))))}), expression(hat(R)(t))),
legend.position='bottomright', legend.inset=0.03,
cex.legend=1, ...)

```

Arguments

X	a matrix of k columns in which each column is the vector of (bio)marker values corresponding to each sample.
D	the vector of response values.
method	the method used to approximate the statistic distribution. One of "general.bootstrap" (<i>Martinez-Cambor et al. (2012)</i>), "permutation" (<i>Venkatraman et al. (1996)</i>) or "auc" (<i>DeLong et al. (1988)</i>).
statistic	the statistic used to compare the curves. One of "KS" (Kolmogorov-Smirnov criteria), "L1" (L_1 -measure), "L2" (L_2 -measure), "CR" (Cramer-von Mises), "other" (another statistic defined by the FUN.dist input parameter), "VK" (Venkatraman) or "AUC" (area under the curve).
FUN.dist	the distance considered as a function of one variable. If statistic="other" the statistic considered is $\sum_{i=1}^k \text{FUN.dist}(\sqrt{n_1}(\hat{R}_i(t) - \hat{R}(t)))$ where n_1 is the number of cases, $\hat{R}_i(t)$ is the ROC curve estimate from the i-th sample and $\hat{R}(t) := k^{-1} \sum_{i=1}^k \hat{R}_i(t)$.
side	type of ROC curve. One of "right" or "left". If method="VK" only right-sided could be considered.
Ni	number of subintervals of the unit interval (FPR values) considered to calculate the curve. Default: 1000.
B	number of bootstrap samples if method="general.bootstrap". Default: 500.
perm	number of permutations if method="permutation". Default: 500.
seed	seed considered to generate the permutations (for reproducibility). Default: 123.
h.fun	a function defining the bandwidth calculus used to generate the bootstrap samples if method="general.bootstrap". It has two arguments: the first one referred to the H value and the second one, x, referred to the sample. Default: <code>function(H,x){H*sd(x)*length(x)^{-1/3}}</code> .
H	the value used to compute h.fun, that is, the bandwidth. Default: 1.
plot.roc	if TRUE, a plot including ROC curve estimates for the k samples and the mean of all of them is displayed.
type	what type of plot should be drawn.
lwd	the line width to be used for mean ROC curve estimate.
lwd.curves	a vector with the line widths to be used for ROC curve estimates of each sample.

lty	the line type to be used for mean ROC curve estimate.
lty.curves	a vector with the line types to be used for ROC curve estimates of each sample.
col	the color to be used for mean ROC curve estimate.
col.curves	a vector with the colors to be used for ROC curve estimates of each sample.
cex.lab	the magnification to be used for x and y labels relative to the current setting of cex.
legend	a character or expression vector to appear in the legend.
legend.position, legend.inset, cex.legend	the position of the legend, the inset distance from the margins as a fraction of the plot region when legend is placed and the character expansion factor relative to current par("cex"), respectively.
...	another graphical parameters to be passed.

Details

First of all, the data introduced is checked and those subjects with some missing information (marker or response value(s)) are removed. Data from a paired design should have the same length along the samples. If this is not fulfilled the code will not run and an error will be showed.

If the Venkatraman statistic is chosen in order to compare left-sided ROC curves, an error will be displayed and it will not work. The Venkatraman methodology is just implemented for right-sided ROC curves. Furthermore, for this statistics, method="permutation" is automatically assigned.

The statistic is defined by $\sum_{i=1}^k \text{FUN.dist}(\sqrt{n_1} \cdot (\hat{R}_i(t) - \hat{R}(t)))$ where FUN.dist stands by the distance function, n_1 is the number of cases, $\hat{R}_i(t)$ is the ROC curve estimate from the i -th sample and $\hat{R}(t) := k^{-1} \sum_{i=1}^k \hat{R}_i(t)$.

The statistics implemented are defined by the following FUN.dist functions:

- statistic="KS":
FUN.dist(g) = max(abs(g))
- statistic="L1":
FUN.dist(g) = mean(abs(g))
- statistic="L2":
FUN.dist(g) = mean(g^2)
- statistic="CR":
FUN.dist.CR(g,h) = sum(g[-length(g)]^2*(h[-1]-h[-length(h)]))
Cramer von-Mises statistic is defined by $\sum_{i=1}^k \text{FUN.dist.CR}(\sqrt{n_1} \cdot (\hat{R}_i(t) - \hat{R}(t)), \hat{R}(t))$

In case of statistic="VK" the Venkatraman methodology (see *References* below) is computed to calculate the statistic. If $k > 2$ the statistic value is the sum of statistic values of each pair such that $i < j$.

If method="general.bootstrap" it is necessary to have a bandwidth in order to compute the bootstrap samples from the smoothed (the gaussian kernel is considered) multivariate empirical distribution functions referred to controls and cases. This bandwidth is defined by the h.FUN function whose parameters are a bandwidth constant parameter defined by the user, H, and the sample (cases or controls values of the marker) considered, x.

If `method="auc"`, the methodology proposed by *DeLong et al.* is implemented. This option is slower because of the Mann-Whitney statistic inside requires *number of cases*·*number of controls* comparisons. In this case, `statistic` returns the value of the Mann-Whitney statistic estimate and `test.statistic` the final test statistic estimate (formula (5) in the paper) which follows a chi-square distribution.

Value

<code>n.controls</code>	the number of controls.
<code>n.cases</code>	the number of cases.
<code>controls.k</code>	a matrix whose columns are the controls along the k samples.
<code>cases.k</code>	a matrix whose columns are the cases along the k samples.
<code>statistic</code>	the value of the test statistic.
<code>stat.boot</code>	a vector of statistic values for bootstrap replicates if <code>method="general.bootstrap"</code> .
<code>stat.perm</code>	a vector of statistic values for permutations if <code>method="permutation"</code> .
<code>test.statistic</code>	statistic estimate given in formula (5) of <i>DeLong et al. (1988)</i> (See <i>References</i> below) if <code>method="auc"</code> .
<code>p.value</code>	the p-value for the test.

References

- Venkatraman E.S., Begg C.B., 1996, A distribution-free procedure for comparing receiver operating characteristic curves from a paired experiment, *Biometrika*, **83**(4), 835-848.
- Martinez-Camblor P., Corral, N., 2012, A general bootstrap algorithm for hypothesis testing, *Journal of Statistical Planning and Inference*, **142**, 589-600.
- Martinez-Camblor P., Carleos C., Corral N., 2013, General nonparametric ROC curve comparison, *Journal of the Korean Statistical Society*, **42**(1), 71-81.
- DeLong E.R., DeLong D.M., Clarke-Pearson D.L., 1988, Comparing the areas under two or more correlated receiver operating characteristic curves: a nonparametric approach, *Biometrics*, **44**, 837-845.

Examples

```
n0 <- 45; n1 <- 60
set.seed(123)
D <- c(rep(0,n0), rep(1,n1))

library(mvtnorm)
rho.12 <- 1/4; rho.13 <- 1/4; rho.23 <- 0.5
sd.controls <- c(1,1,1)
sd.cases <- c(1,1,1)
var.controls <- sd.controls%*%t(sd.controls)
var.cases <- sd.cases%*%t(sd.cases)
sigma.controls <- var.controls*matrix(c(1,rho.12,rho.13,rho.12,1,rho.23,rho.13,rho.23,1),3,3)
sigma.cases <- var.cases*matrix(c(1,rho.12,rho.13,rho.12,1,rho.23,rho.13,rho.23,1),3,3)
controls <- rmvnorm(n0, mean=rep(0,3), sigma=sigma.controls)
cases <- rmvnorm(n1, mean=rep(1.19,3), sigma=sigma.cases)
```

```

marker.samples <- rbind(controls,cases)

# Default method: KS statistic proposed in Martinez-Cambor by general bootstrap
output <- compareROCdep(marker.samples, D)

# L1 statistic proposed in Martinez-Cambor by general bootstrap
output1 <- compareROCdep(marker.samples, D, statistic="L1")

# CR statistic proposed in Martinez-Cambor by permutation method
output2 <- compareROCdep(marker.samples, D, method="permutation", statistic="CR")

# Venkatraman statistic
output3 <- compareROCdep(marker.samples, D, statistic="VK")

# DeLong AUC comparison methodology
output4 <- compareROCdep(marker.samples, D, method="auc")

```

compareROCindep

Comparison of k independent ROC curves

Description

This function compares k ROC curves from independent data. Different statistics can be considered in order to perform the comparison: those ones included in *Martinez-Cambor et al. (2011)* based on distances, the *Venkatraman (2000)* methodology for comparing curves for continuous unpaired data and one based in AUC (area under the curve) comparison. See *References* below.

Usage

```

compareROCindep(X, G, D, ...)
## Default S3 method:
compareROCindep(X, G, D, statistic=c("L1","L2","CR","other","VK","AUC"),
  FUN.stat.int=function(roc.i, roc){mean(abs(roc.i - roc))},
  FUN.stat.cons=function(n.cases, n.controls){sqrt(n.cases)},
  side=c("right","left"), Ni=1000, raw=FALSE, perm=500,
  seed=123, plot.roc=TRUE, type='s', lwd=3,
  lwd.curves=rep(2,length(table(G))), lty=1,
  lty.curves=rep(1,length(table(G))), col='black',
  col.curves=rainbow(length(table(G))), cex.lab=1.2,
  legend=c(sapply(1:length(table(G)),function(i){
    eval(bquote(expression(hat(R)[.(i)](t))))},
  expression(hat(R)(t))), legend.position='bottomright',
  legend.inset=0.03, cex.legend=1, ...)

```

Arguments

X vector of (bio)marker values.

G	vector of group identifier values (it should have as levels as independent samples to compare).
D	the vector of response values.
statistic	the statistic used in order to compare the curves. One of "L1" (L_1 -measure), "L2" (L_2 -measure), "CR" (Cramer-von Mises), "other" (another statistic defined by $\sum_{i=1}^k \text{FUN.stat.cons} \cdot \text{FUN.stat.int}$), "VK" (Venkatraman) or "AUC" (area under the curve).
FUN.stat.int	a function of two variables, roc.i and roc standing for ROC curve estimate for the i -th sample and mean ROC curve estimate along the k samples, respectively. This function represents the integral to consider in case of statistic="other".
FUN.stat.cons	a function of two variables, n.cases and n.controls standing for the cases and controls sample size, respectively. This function represents the constant to multiply FUN.stat.int above in case of statistic="other".
side	type of ROC curve. One of "right" or "left". If method="VK" only right-sided could be considered.
Ni	number of subintervals of the unit interval (FPR values) considered to calculate the curve. Default: 1000.
raw	if TRUE, raw data is considered; if FALSE, data is ranked and a method to break ties in the permutations is considered (see <i>Venkatraman (2000)</i> in <i>References</i>). Default: FALSE.
perm	number of permutations. Default: 500.
seed	seed considered to generate the permutations (for reproducibility). Default: 123.
plot.roc	if TRUE, a plot including ROC curve estimates for the k samples and the mean of all of them is displayed.
type	what type of plot should be drawn.
lwd	the line width to be used for mean ROC curve estimate.
lwd.curves	a vector with the line widths to be used for ROC curve estimates of each sample.
lty	the line type to be used for mean ROC curve estimate.
lty.curves	a vector with the line types to be used for ROC curve estimates of each sample.
col	the color to be used for mean ROC curve estimate.
col.curves	a vector with the colors to be used for ROC curve estimates of each sample.
cex.lab	the magnification to be used for x and y labels relative to the current setting of cex.
legend	a character or expression vector to appear in the legend.
legend.position, legend.inset, cex.legend	the position of the legend, the inset distance from the margins as a fraction of the plot region when legend is placed, and the character expansion factor relative to current par("cex"), respectively.
...	another graphical parameters to be passed.

Details

If the Venkatraman statistic is chosen in order to compare left-sided ROC curves, an error will be displayed and it will not work. The Venkatraman methodology is just implemented for right-sided ROC curves.

If `raw=FALSE` the data will be ranked in each sample using the rank function with `ties.method='first'` option. Furthermore, the permutation samples possible ties will be broken using `ties.method='random'` option.

The statistic is defined by $\sum_{i=1}^k \text{statistic.cons} \cdot \text{statistic.int}$ where `statistic.cons` = `FUN.stat.cons('number of cases in the i-th sample', 'number of controls in the i-th sample')` and `statistic.int` = `FUN.stat.int('ROC curve estimate from the i-th sample', 'mean ROC curve estimate along the k samples')`. It is usual to consider the function `FUN.stat.int` as an integral of a distance between $\hat{R}_i(t)$ and $\hat{R}(t)$ where $\hat{R}(t) := k^{-1} \sum_{i=1}^k \hat{R}_i(t)$.

The statistics implemented are defined by the following `FUN.stat.cons` and `FUN.stat.int` functions:

- `statistic="L1"`:
`FUN.stat.int(roc.i, roc) = mean(abs(roc.i - roc))`
`FUN.stat.cons(n.cases, n.controls) = sqrt(n.cases)`
- `statistic="L2"`:
`FUN.stat.int(roc.i, roc) = mean((roc.i - roc)^2)`
`FUN.stat.cons(n.cases, n.controls) = n.cases`
- `statistic="CR"`:
`FUN.stat.int(roc.i, roc) = mean((roc.i[seq(2, 2*Ni+1, 2)] - roc[seq(2, 2*Ni+1, 2)])^2 * (roc[seq(3, 2*Ni+1, 2)] - roc[seq(1, 2*Ni-1, 2)]))`
`FUN.stat.cons(n.cases, n.controls) = n.cases`
 In order to use this statistic, the ROC curves have been estimated in a grid with $2 \cdot Ni$ subintervals of the unit interval.

The permutation method proposed in *Venkatraman (2000)* is used in order to generate the perm samples in all methodologies (i.e., any statistic).

In case of `statistic="VK"` the Venkatraman methodology (see *References* below) is computed to calculate the statistic. If $k > 2$ the statistic value is the sum of the statistic values of each pair such that $i < j$.

In case of `statistic="AUC"`, the statistic considered is $k^{-1} \sum_{i=1}^k |\widehat{AUC}'_i - \widehat{AUC}'|$ where \widehat{AUC}' is the mean of \widehat{AUC}'_i along the k samples.

Value

<code>n.controls</code>	vector of number of controls in each sample.
<code>n.cases</code>	vector of number of cases in each sample.
<code>controls.k</code>	a vector of all controls along the k samples, ordered by sample.
<code>cases.k</code>	a vector of all cases along the k samples, ordered by sample.
<code>statistic</code>	the value of the test statistic.
<code>stat.perm</code>	a vector of statistic values for permutations.
<code>p.value</code>	the p-value for the test.

References

Venkatraman E.S., 2000, A permutation test to compare receiver operating characteristic curves, *Biometrics*, **56**, 1134-1138.

Martinez-Cambor P., Carleos C., Corral N., 2011, Powerful nonparametric statistics to compare k independent ROC curves, *Journal of Applied Statistics*, **38**(7), 1317-1332.

Examples

```
set.seed(123)
X1 <- c(rnorm(45), rnorm(30,2,1.5))
D1 <- c(rep(0,45), rep(1,30))
X2 <- c(rnorm(45), rnorm(38,3,1.5))
D2 <- c(rep(0,45), rep(1,38))
X3 <- c(rnorm(30), rnorm(42,3,1))
D3 <- c(rep(0,30), rep(1,42))
X <- c(X1, X2, X3)
D <- c(D1, D2, D3)
G <- c(rep(1,75), rep(2,83), rep(3,72))

# Default method: L1 statistic proposed in Martinez-Cambor
output <- compareROCindep(X, G, D)

# Venkatraman statistic
output1 <- compareROCindep(X, G, D, statistic="VK")

# DeLong AUC comparison methodology
output2 <- compareROCindep(X, G, D, statistic="AUC")
```

gROC

ROC curve estimation (generalization included)

Description

This function estimates the Receiver Operating Characteristic curve and returns a 'groc' object. This object can be printed or plotted. It is the main function of the package and it includes the ROC curve generalization for non-monotone relationships (see *References* below).

Usage

```
gROC(X, D, ...)
## Default S3 method:
gROC(X, D, side=c("right", "left", "both", "both2", "auto"),
      Ni = NULL, plot.roc = FALSE, plot.density = FALSE,
      pval.auc = FALSE, B = 500, ...)
```

Arguments

X	vector of (bio)marker values.
D	vector of response values.
side	type of ROC curve. One of "right", "left", "both", "both2" or "auto". If the user does not specify the method, "right" is considered, i.e., the right-sided ROC curve is estimated. If "auto", one of "right" or "left" is assigned automatically according to the Wilcoxon statistic. If the estimate is lower than $\frac{\text{number of controls} \times \text{number of cases}}{2}$, the right-sided ROC curve is computed, otherwise the left one is considered. If "both", the gROC curve is estimated in the usual way (controls inside the interval) and if "both2", the opposite direction is considered.
Ni	number of subintervals of the unit interval (FPR values) considered to calculate the curve. Default: NULL (in this case the fast algorithm considering as many FPR values as number of controls is considered).
plot.roc	if TRUE, ROC curve estimate considered is displayed.
plot.density	if TRUE, density estimates for the (bio)marker in the positive and negative subjects are displayed in the same plot.
pval.auc	if TRUE, a permutation test to test if the AUC differs from 0.5 is performed. Available if Ni is NULL. Default: FALSE.
B	number of permutations used for testing. Default: 500.
...	additional arguments for gROC. Ignored.

Details

First of all, the data introduced is checked by the ROCcheck function.

If side is not specified, one of "right" or "left" options is considered according to the comparison of the Wilcoxon test estimate and $\frac{\text{number of controls} \times \text{number of cases}}{2}$. In this case, Wilcoxon rank test is performed in order to test the alternative hypothesis $\text{median}(\text{controls}) < \text{median}(\text{cases})$ and the resulting p-value is shown but side selection is just based in the aforementioned comparison, without taking into account the p-value of the test.

If Ni is NULL, the general ROC curve, $R_g(\cdot)$ is estimated considering every different pair $(t, \gamma \cdot t)$ available on data. Otherwise it is estimated using $\{0, 1e-3, 2e-3, \dots, 1\}$ as a grid for γ in the unit interval.

If both plot.density and plot.roc are TRUE they are displayed in the same window.

Value

A list of class "groc" with the following content:

levels	levels in D. The two first ones are the labels of D considered controls and cases, respectively.
controls	marker values for controls.
cases	marker values for cases.

side	direction of the comparison between controls and cases. One of "right", "left", "both" or "both2". If side is auto in input arguments, the output will be the one considered by the function (according to the comparison specified above).
pvalue.wilcox	p-value of Wilcoxon test performed to compare cases and controls. Alternative hypothesis: $median(controls) < median(cases)$.
t	vector of values of t: $\{0, 1/N_i, 2/N_i, \dots, 1\}$.
roc	vector of values of $R(t)$ for each t .
auc	area under the ROC curve estimate by trapezoidal rule.
pval.auc	p-value of the permutation test over the AUC.
Paucs	different permutation AUCs displayed if the hypothesis test is performed.
points.coordinates	coordinates of the points (FPR, TPR) where the ROC curve estimate has a step in case of right or left-sided curves. The first column corresponds to the values of the points vector providing these coordinates.
pairpoints.coordinates	coordinates of the points (FPR, TPR) where the ROC curve estimate has a step in case of general curves. The first and second columns correspond to the values of the pairpoints matrix providing these coordinates.
param	a logical value indicating if the estimation procedure whether parametric or not.
Ni	number of subintervals of the unit interval considered to build the curve.
points	if $N_i \neq \text{NULL}$, vector of cut-off points of the (bio)marker considered to estimate left or right-sided ROC curves. It corresponds to the sorted marker-values, including $-\infty$ and ∞ .
pairpoints	if $N_i \neq \text{NULL}$, matrix whose rows correspond to each pair of cut-off points (x_l, x_u) such that $x_l < x_u$ of the (bio)marker considered to estimate general ROC curves.
specificities	if $N_i \neq \text{NULL}$, vector of specificities associated to points if left or right-sided ROC curves are considered and those ones corresponding to pairpoints in case of general ROC curves. It could contain repeated values in both cases.
sensitivities	if $N_i \neq \text{NULL}$, vector of sensitivities associated to points if left or right-sided ROC curves are considered and those ones corresponding to pairpoints in case of general ROC curves. It could contain repeated values in both cases.
coordinates	if $N_i \neq \text{NULL}$, matrix whose first column is constituted by the unique values of FPR (1-specificities) and its second column is formed by the corresponding values of TPR (sensitivities). In the case of general ROC curves, several TPR values could be associated with the same FPR value.
index	if $N_i \neq \text{NULL}$, in general ROC curves, a vector displaying which row(s) of the pairpoints matrix correspond(s) to pairpoints.coordinates.

References

Martinez-Cambor P., Corral N., Rey C., Pascual J., Cernuda-Morollon E., 2014, ROC curve generalization for non-monotone relationships, *Statistical Methods in Medical Research*, **26**(1), 113-123.

Examples

```
# Basic example (side="auto") -> Output side is "right"
set.seed(123)
X <- c(rnorm(45), rnorm(30,2,1.5))
D <- c(rep(0,45), rep(1,30))
gROC(X,D)

# Basic example (side="auto") -> Output side is "left"
X <- c(rnorm(45), rnorm(30,-2,1.5))
D <- c(rep(0,45), rep(1,30))
gROC(X,D)

# General ROC curve example
X <- c(rnorm(45), rnorm(30,1,4))
D <- c(rep(0,45), rep(1,30))
gROC(X, D, side="both")

# Plot density estimates and ROC curve in the same plot
X <- c(rnorm(45), rnorm(30,2,1.5))
D <- c(rep(0,45), rep(1,30))
gROC(X, D, plot.roc=TRUE, plot.density=TRUE)
```

interleukin6

Interleukin 6 (IL6) data

Description

This dataset includes the true-positives (TP), false-positives (FP), true-negatives (TN) and false-negatives (FN) reported by 9 different papers which study the use of the Interleukin 6 (IL6) as a marker for the early detection of neonatal sepsis.

Usage

```
interleukin6
```

Format

A data frame with 19 observations of the following 5 variables.

Author a vector assigning different numbers to each paper

TP vector of true positives

FP vector of false positives

FN vector of false negatives

TN vector of true negatives

Details

In those papers providing more than one pair of Sensitivity-Specificity all of them are collected.

References

Martinez-Camblor P., 2017, Fully non-parametric receiver operating characteristic curve estimation for random-effects meta-analysis, *Statistical Methods in Medical Research*, **26**(1), 5-20.

Examples

```
# Load the dataset
data(interleukin6)

# Plot pairs (FPR, TPR) for each Author

attach(interleukin6)

TPR <- TP/(TP+FN)
FPR <- FP/(FP+TN)
plot(FPR, TPR, xlim=c(0,1), ylim=c(0,1), lwd=10, pch=1, col='gray', xlab="False-Positive Rate",
      ylab="True-Positive Rate", main=paste("ROC curve interpolation"))

S <- unique(Author)
ind <- order(Author, FPR, TPR)
ord.data <- cbind(Author[ind], FPR[ind], TPR[ind])
roc.j <- sapply(S, function(j){
  lines(c(0,ord.data[Author==j,2],1), c(0,ord.data[Author==j,3],1), col='gray')}})
for(i in 1:19){text(ord.data[i,2],ord.data[i,3],ord.data[i,1],cex=0.5)}
```

metaROC

Non-parametric ROC curve estimate for meta-analysis

Description

This function performs meta-analytic studies of diagnostic tests for both the fixed and random-effects models. In particular it reports a fully non-parametric ROC curve estimate when data come from a meta-analysis study using the information of all cut-off points available in the selected original studies. The approach considered is the one proposed by *Martinez-Camblor et al. (2017)* based on weighting each individual interpolated ROC curve. See *References* below.

Usage

```
metaROC(data, ...)
## Default S3 method:
metaROC(data, Ni=1000, model=c("fixed-effects","random-effects"),
        plot.Author=FALSE, plot.bands=TRUE, plot.inter.var=FALSE,
        cex.Author=0.7, lwd.Author=12, col.curve='blue',
        col.bands='light blue', alpha.trans=0.5, col.border='blue', ...)
```

Arguments

<code>data</code>	a data frame containing at least the following variables (with these names): <ul style="list-style-type: none"> • Author: a vector assigning different numbers to each paper/author. • TP: true positives. • FP: false positives. • TN: true negatives. • FN: false negatives.
<code>Ni</code>	number of points of the unit interval (FPR values) considered to calculate the curve. Default: 1000.
<code>model</code>	the meta-analysis model used to estimate the ROC curve. One of "fixed-effects" (it only considers the within-study variability) or "random-effects" (it takes into account the variability between the studies).
<code>plot.Author</code>	if TRUE, a plot including ROC curve estimates (by linear interpolation) for each paper under study is displayed.
<code>plot.bands</code>	if TRUE, confidence interval estimate for the curve is added to the plot of the ROC curve estimate.
<code>plot.inter.var</code>	if TRUE, a plot including inter-study variability estimate is displayed on an additional window.
<code>cex.Author</code>	the magnification to be used to display the paper/author points labels relative to the current setting of <code>cex</code> .
<code>lwd.Author</code>	the size to be used for the paper/author points.
<code>col.curve</code>	the color to be used for the (summary) ROC curve estimate. Default: blue.
<code>col.bands</code>	the color to be used for the confidence interval of ROC curve estimate. Default: light blue.
<code>alpha.trans</code>	proportion of opacity to be used for the confidence interval of ROC curve estimate. A number in the unit interval where 0 means transparent. Default: 0.5.
<code>col.border</code>	the color to be used for the border of confidence interval of ROC curve estimate. Default: blue.
<code>...</code>	another graphical parameters to be passed.

Details

The slight modification considered to ensure the monotonicity of the summary ROC curve estimate is the following $sRA(t) = \max(\sup_{z \in [0, t]} sRA(z), RA(t))$.

Some basic information about the model used and the results obtained are printed.

Value

<code>data</code>	the data-frame considered ordered by Author-FPR-TPR and including the following variables: <ul style="list-style-type: none"> • n: positive subjects sample size. • m: negative subjects sample size. • FPR: false positive rate.
-------------------	--

- TPR: true positive rate.

t	values of the unit interval (FPR values) considered to calculate the curve.
model	the meta-analysis model used to estimate the ROC curve. One of "fixed-effects" (it only considers the within-study variability) or "random-effects" (it takes into account the variability between the studies).
sRA	non-parametric summary ROC curve estimate following the model considered with a slight modification to ensure the monotonicity. This is the one reported in graphics.
RA	non-parametric summary ROC curve estimate following the model without modifications.
se.RA	standard-error of summary ROC curve estimate.
area	area under the summary ROC curve estimate by trapezoidal rule.
youden.index	the optimal specificity and sensitivity (in the Youden index sense).
roc.j	a matrix whose column j contains the estimated ROC curve for the j-th study in each point t considered.
w.j	a matrix whose column j contains the weights in fixed-effects model for the j-th study in each point t considered.
w.j.rem	a matrix whose column j contains the weights in random-effects model for the j-th study in each point t considered.
inter.var	inter-study variability estimate in each point t considered. Only computed if model is "random-effects".

References

Martinez-Camblor P., 2017, Fully non-parametric receiver operating characteristic curve estimation for random-effects meta-analysis, *Statistical Methods in Medical Research*, **26**(1), 5-20.

Examples

```
data(interleukin6)

# Fixed-effects meta-analysis showing linear interpolations of the papers considered in the graphic
output1 <- metaROC(interleukin6, plot.Author=TRUE)

# Random-effects meta-analysis displaying also a window with a plot of the inter-study
# variability estimate
output2 <- metaROC(interleukin6, model="random-effects", plot.Author=TRUE)
```

 plot

Plot an ROC curve object

Description

This function plots a 'groc', 'rocbands' or 'cdroc' object.

Usage

```
## S3 method for class 'groc'
plot(x, lwd = 2, xlab = "False-Positive Rate",
     ylab = "True-Positive Rate", main = "ROC curve", ...)
## S3 method for class 'rocbands'
plot(x, type='s', lwd=2, xlim=c(0,1), ylim=c(0,1),
     xlab="False-Positive Rate", ylab="True-Positive Rate",
     main=paste("ROC curve \n (", obj$method, " confidence bands)", sep=""),
     col='aquamarine3', col.inside="azure2", col.frontier="azure3",
     lwd.frontier=2, ...)
## S3 method for class 'cdroc'
plot(x, type='s', lwd=3, xlab='1 - Specificity',
     ylab='Sensitivity', xaxs='i', yaxs='i',
     main=paste("ROC curve at time", obj$predict.time), ...)
```

Arguments

x	a 'groc', 'rocbands' or 'cdroc' object from the gROC, ROCbands or cdROC respectively.
type	what type of plot should be drawn.
lwd	the line width to be used for ROC curve estimate, a positive number. See <code>par</code> .
col	the color to be used for ROC curve estimate. See <code>par</code> .
lwd.frontier	the line width to be used for ROC curve confidence bands estimate.
col.inside, col.frontier	the color to be used for ROC curve confidence bands estimate (<code>col.frontier</code>) and for the area inside (<code>col.inside</code>).
xlim, ylim	numeric vectors of length 2, giving the x and y coordinates ranges. See <code>plot.window</code> .
xlab, ylab	a title for the x and y axis, respectively. See <code>title</code> .
xaxs, yaxs	the style of axis interval calculation to be used for the x and y axis, respectively. See <code>par</code> .
main	an overall title for the plot. See <code>title</code> .
...	further arguments to be passed to methods, such as graphical parameters. See <code>par</code> .

Value

These functions return a plot of the object they were passed.

Examples

```
# Data generation
set.seed(123)
X <- c(rnorm(45), rnorm(30,2,1.5))
D <- c(rep(0,45), rep(1,30))

# Plot an ROC curve
grocobj <- gROC(X,D)
plot(grocobj)

# Plot ROC curve confidence bands
rocbandsobj <- ROCbands(grocobj)
plot(rocbandsobj)

# Plot cumulative/dynamic ROC curve
set.seed(123)
stime <- rchisq(50,3)
status <- sample(c(rep(1,40), rep(0,10)))
marker <- max(stime) - stime + rnorm(50,0,2)
cdrocobj <- cdROC(stime, status, marker, 2.8, ci=TRUE)
plot(cdrocobj)
```

print

Print an ROC curve object

Description

This function prints a 'groc', 'rocbands' or 'cdroc' object.

Usage

```
## S3 method for class 'groc'
print(x, ...)
## S3 method for class 'rocbands'
print(x, ...)
## S3 method for class 'cdroc'
print(x, ...)
```

Arguments

x a 'groc', 'rocbands' or 'cdroc' object from the gROC, ROCbands or cdROC respectively.

... further arguments to be passed to other methods. Ignored.

Value

These functions return information about the object they were passed.

See Also

[gROC](#), [ROCbands](#), [cdROC](#)

Examples

```
# Data generation
set.seed(123)
X <- c(rnorm(45), rnorm(30,2,1.5))
D <- c(rep(0,45), rep(1,30))

# Print a groc object
grocobj <- gROC(X,D)
print(grocobj)

# Print an rocbands object
grocobj <- ROCbands(grocobj)
print(grocobj)

# Print a cdroc object
set.seed(123)
stime <- rchisq(50,3)
status <- sample(c(rep(1,45), rep(0,5)))
marker <- max(stime) - stime + rnorm(50,0,2)
cdrocobj <- cdROC(stime, status, marker, 3, ci=TRUE)
print(cdrocobj)
```

ROCbands

Confidence bands for ROC curves

Description

This function computes and plots confidence bands for ROC curves (both left/right-sided and general one) using three different procedures. Particularly, one parametric approach assuming the binormal model (*Demidenko*) and two non-parametric techniques (*Jensen et al.* and *Martinez-Cambor et al.*). See *References* below.

Usage

```
ROCbands(groc, ...)
## Default S3 method:
ROCbands(groc, method = c("PSN", "JMS", "DEK"), conf.level = 0.95,
          B = 500, bootstrap.bar = TRUE, alpha1 = NULL, s = 1, a.J = NULL, b.J = NULL,
          plot.bands = FALSE, plot.var = FALSE, seed = 123, ...)
```

Arguments

`groc` a 'groc' object from the `gROC` function.

`method` method used to compute the confidence bands. One of "PSN" (*Martinez-Cambor et al.*), "JMS" (*Jensen et al.*) or "DEK" (*Demidenko*).

conf.level	the width of the confidence band as a number in (0,1). Default: 0.95, resulting in a 95% confidence band.
B	number of bootstrap replicates. Default: 500 (only used in "PSN" and "JMS" methods).
bootstrap.bar	if TRUE, a bar showing bootstrap replication progress is displayed.
alpha1	α_1 in "PSN" approach a number in (0,1) affecting the width between the lower band and the ROC curve estimate. Default: NULL, the one which minimizes the theoretical area between lower and upper bands is considered.
s	scale parameter used to compute the smoothed kernel distribution functions in "PSN" method. The bandwidth $h = s \cdot \min(m, n)^{1/5} \cdot \hat{\sigma}$ where m and n stand by the number of controls and cases, respectively, is considered. Default: 1.
a. J, b. J	extremes of interval in (0,1) in which compute the regional confidence bands by "JMS" methodology. Default: (1/Ni, 1 - 1/Ni).
plot.bands	if TRUE, confidence bands at level conf.level are displayed.
plot.var	if TRUE, a plot of $\sigma_n^{*,1}(t)$ with t in [0,1] (if "PSN" method is selected) or $Var(\Psi(p))$ with p in (a. J, b. J) (if "JMS" method is selected) is displayed.
seed	seed used to compute the bootstrap controls and cases samples in "PSN" method or Brownian Bridges in "JMS" method.
...	additional arguments for ROCbands. Ignored.

Details

- *Martinez-Camblor et al. methodology* - "PSN" method

The theoretical area is computed as $(c_{\alpha_1} - c_{\alpha_2})n^{-1/2} \int \sigma_n^*(t)dt$ where $\sigma_n^*(t)$ is the standard deviation estimate of $\sqrt{n}[\hat{R}(\omega, \cdot) - R(\cdot)]$ and n is the cases sample size.

Due to computation can take some time depending on the number of bootstrap replicates considered, a progress bar is shown.

Confidence bands are truncated in the following way: on one hand, if the lower band is lower than 0 or higher than 0.95 it is forced to be 0 or 0.95, respectively; on the other hand, if the upper band is higher than 1 or lower than 0.05 it is forced to be 1 or 0.05, respectively.

- *Jensen et al. methodology* - "JMS" method

$K_{a,b}^\alpha$ denote the upper $\alpha/2$ -quantile of the distribution of $\sup_{a \leq p \leq b} \frac{|\Psi(p)|}{\sqrt{Var\Psi(p)}}$ where (a, b) is the interval in which the regional confidence bands are calculated and $\Psi(\cdot)$ is the limiting process of the stochastic process $\Delta_N = \sqrt{N}[\hat{R}(\omega, \cdot) - R(\cdot)]$ with N being the total sample size.

Extremes of the interval (a. J, b. J) used in order to display the regional confidence bands must be divisors of Ni in the interval [0, 1].

Confidence bands are truncated in a similar way as in "PSN" method in order not to have bands lower than 0 or higher than 1.

- *Demidenko methodology* - "DEK" method

Demidenko ROC curve estimate does not correspond to the empirical one due to the fact that the (bio)marker values in controls and cases are supposed to come from a normal distribution is exploited.

Value

A list of class 'rocbands' with the following content:

method	method used to compute the confidence bands. One of "PSN" (Martinez-Camblor et al.), "JMS" (Jensen et al.) or "DEK" (Demidenko).
conf.level	the width of the confidence band as a number in (0,1).
B	number of bootstrap replicates used in "PSN" and "JMS" methods.
L, U	vectors containing the values of lower and upper bands, respectively, for each $t \in \{0, 1/N_i, 2/N_i, \dots, 1\}$. In case of "JMS" method p is considered as t .
practical.area	area between lower and upper bands (L and U) computed by trapezoidal rule.
Ni	number of subintervals of the unit interval considered to build the curve.
ROC.t	vector of values of $R(t)$ for each $t \in \{0, 1/N_i, 2/N_i, \dots, 1\}$.

If the method is "PSN":

s	scale parameter used to compute the smoothed kernel distribution functions.
alpha1, alpha2	if the alpha1 input argument is not specified, α_1 and α_2 values which minimize area between bands are automatically computed. If alpha1 is chosen by the user, alpha2 is computed by $\alpha_2 = (1 - \text{conf.level}) - \alpha_1$.
fixed.alpha1	if TRUE, alpha1 has been fixed by the user.
c1, c2	c_{α_1} and c_{α_2} resulting from the algorithm to compute confidence bands.
ROC.B	matrix of size N_i+1, B whose columns contain the ROC curve estimate for each bootstrap sample.
sd.PSN	vector $\sigma_n^*(t)$ which is the estimate of the standard deviation of the empirical process considered.
theoretical.area	theoretical area between confidence bands by trapezoidal rule.

If the method is "JMS":

a.J, b.J	extremes of the interval in which the regional confidence bands have been computed.
p	vector of FPR points considered in the interval (a.J, b.J).
smoothROC.p	smooth ROC curve estimate for each value of p.
K.alpha	value of $K_{a,b}^\alpha$ computed to calculate confidence bands (see Details above).
var.JMS	value of $Var(\Psi(p))$ estimated from the formula given by Hsieh and Turnbull (see <i>Jensen et al.</i> in <i>References</i>).

If the method is "DEK":

DEK.fpr, DEK.tpr	values of FPR and TPR computed to calculate the Demidenko confidence bands taking into account that it is a binormal technique.
------------------	---

Note

Brownian bridges needed to estimate $\Psi(\cdot)$ in "JMS" method are computed using the BBridge function in the sde package.

It should be noted that both the "PSN" and "JMS" methods are non-parametric, while the "DEK" approach is designed assuming the binormal model, so it is not convenient to use this method when distribution assumptions are not fulfilled. Furthermore, both the "JMS" and "DEK" methodologies are implemented just for the right-sided ROC curve. If side is left or both only the "PSN" method provides confidence bands.

References

Martinez-Cambor P., Perez-Fernandez S., Corral N., 2016, Efficient nonparametric confidence bands for receiver operating-characteristic curves, *Statistical Methods in Medical Research*, DOI: 10.1177/0962280216672490.

Jensen K., Muller H-H., Schafer H., 2000, Regional confidence bands for ROC curves, *Statistical in Medicine*, **19**, 493-509.

Demidenko E., 2012, Confidence intervals and bands for the binormal ROC curve, *Journal of Applied Statistics*, **39**(1), 67-79.

Examples

```
# Basic example
set.seed(123)
X <- c(rnorm(45), rnorm(30,2,1.5))
D <- c(rep(0,45), rep(1,30))
groc.obj <- gROC(X,D)

# PSN confidence bands with conf.level=0.95
ROCbands(groc.obj)
# Plot standard deviation estimate of the curve and confidence bands in the same window
ROCbands(groc.obj, plot.bands=TRUE, plot.var=TRUE)
# PSN confidence bands with alpha1 fixed (alpha1=0.025)
ROCbands(groc.obj, alpha1=0.025)

# JMS confidence bands in (0.2,0.7) interval
ROCbands(groc.obj, method="JMS", a.J=0.2, b.J=0.7)
# Plot variance estimate of the curve and confidence bands in the same window
ROCbands(groc.obj, method="JMS", a.J=0.2, b.J=0.7, plot.bands=TRUE, plot.var=TRUE)

# DEK confidence bands with conf.level=0.99
ROCbands(groc.obj, method="DEK", conf.level=0.99)
```

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